

# Bounded Independence vs. Moduli\*

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## Abstract

Let  $k = k(n)$  be the largest integer such that there exists a  $k$ -wise uniform distribution over  $\{0, 1\}^n$  that is supported on the set  $S_m := \{x \in \{0, 1\}^n : \sum_i x_i \equiv 0 \pmod{m}\}$ , where  $m$  is any integer. We show that  $\Omega(n/m^2 \log m) \leq k \leq 2n/m + 2$ . For  $k = O(n/m)$  we also show that any  $k$ -wise uniform distribution puts probability mass at most  $1/m + 1/100$  over  $S_m$ . For any fixed odd  $m$  there is  $k \geq (1 - \Omega(1))n$  such that any  $k$ -wise uniform distribution lands in  $S_m$  with probability exponentially close to  $|S_m|/2^n$ ; and this result is false for any even  $m$ .

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## 1 Introduction and our results

A distribution on  $\{0, 1\}^n$  is  $k$ -wise uniform if any  $k$  bits are uniform in  $\{0, 1\}^k$ . Researchers have analyzed various classes of tests that cannot distinguish distributions with  $k$ -wise uniformity from uniform. Such tests include (combinatorial) rectangles [8] (cf. [4]), bounded-depth circuits [1, 12, 2, 13], and halfspaces [6, 9, 7], to name a few. We say that such tests are *fooled* by distributions with bounded independence.

In this work we consider the mod  $m$  tests, defined next.

► **Definition 1.** For an input length  $n$ , and an integer  $m$ , we define the set  $S_m := \{x \in \{0, 1\}^n : \sum_i x_i \equiv 0 \pmod{m}\}$ .

These tests have been intensely studied at least since circuit complexity theory hit the wall of gates computing mod  $m$  for composite  $m$  in the 80's. However, the effect of bounded independence on mod  $m$  tests does not seem to have been known before this paper.

Our first main result is that there exist distributions with linear uniformity that are supported on  $S_m$ .

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► **Theorem 2.** *There exists a  $c > 0$  such that the following holds.*

*For every integer  $m \geq 2$ , there exists a  $k \geq cn/m^2 \log m$  and a  $k$ -wise uniform distribution over  $\{0, 1\}^n$  that is supported on  $S_m$ .*

This proves a conjecture in [10] where this question is also raised. Their motivation was a study of the “mod 3” dimension of  $k$ -wise uniform distributions, started in [11], which is the dimension of the space spanned by the support of the distribution over  $\text{GF}(3)$ . [10] shows that  $k = 100 \log n$ -wise uniformity with dimension  $\leq n^{0.49}$  would have applications to pseudorandomness. It also exhibits a distribution with dimension  $n^{0.72}$  and uniformity  $k = 2$ . Theorem 2 yields a distribution with dimension  $n - 1$  and  $\Omega(n)$ -wise uniformity.

We then prove three results, summarized in the next theorem, that show that  $k$ -wise uniformity does fool mod  $m$  when  $k$  is large. (1) shows that the largest possible value of  $k$  in Theorem 2 is  $k \leq 2(n + 1)/m + 2 \leq (1 - \Omega(1))n$ . (2) shows that when  $k$  is larger than  $(1 - \gamma)n$  for a constant  $\gamma$  depending only on  $m$  then  $k$ -wise uniformity fools  $S_m$  with exponentially small error when  $m$  is odd. The proof of (2) however does not carry to the setting of  $k < n/2$ , for any  $m$ . So we establish (3) which gives a worse error bound but allows for  $k$  to become smaller for larger  $m$ , specifically  $k = O(n/m)$  for constant error. The error bound in (3) and the density of  $S_m$  are such that (3) only provides a meaningful upper bound on the probability that the  $k$ -wise uniform distribution lands in  $S_m$ , but not a lower bound. In fact, we conjecture that no lower bound is possible in the sense that there is  $c > 0$  such that for every  $m$  there is a  $cn$ -wise uniform distribution supported on the complement of  $S_m$ .

The combination of (2) and (3) implies that for  $k = \min\{O(n/m), (1 - \Omega(1))n\}$  any  $k$ -wise uniform distribution puts probability mass at most  $1/m + 1/100$  over  $S_m$  for odd  $m$ .

► **Theorem 3.** *Let  $m$  be an integer.*

- (1) *For  $k \geq 2n/m + 2$ , a  $k$ -wise uniform distribution over  $\{0, 1\}^n$  cannot be supported on  $S_m$ .*
- (2) *Suppose  $m$  is odd, then there is a  $\gamma > 0$  depending only on  $m$  such that for any  $(1 - \gamma)n$ -wise uniform distribution  $D$  over  $\{0, 1\}^n$ ,  $|\Pr[D \in S_m] - |S_m|/2^n| \leq 2^{-\gamma n}$ .*
- (3) *There exists a universal constant  $c$  such that for every  $\varepsilon > 0$ ,  $n \geq cm^2 \log(m/\varepsilon)$ , and any  $c(n/m)(1/\varepsilon)^2$ -wise uniform distribution  $D$  over  $\{0, 1\}^n$ ,  $\Pr[D \in S_m] \leq |S_m|/2^n + \varepsilon$ .*

In our results the sum  $s$  of  $n$  bits  $x_i \in \{0, 1\}$  is constrained to be divisible by  $m$ . This setting was chosen for convenience, but our techniques apply in greater generality. For example we obtain the same results if we instead constrain  $s$  to be  $c \pmod m$  for any fixed  $c$ .

We also note that (2) is false for any even  $m$  because the uniform distribution on  $S_2$  has uniformity  $k = n - 1$  but puts about  $2/m$  mass on  $S_m$ , a set which as we shall see later (cf. Remark 7) has density about  $1/m$ .

## Organization

Theorem 3 is a little easier to prove than Theorem 2, but uses overlapping lemmas. So we start by proving Theorem 3 in Section 2. Then in Section 3 we prove Theorem 2.

## 2 Proof of Theorem 3

In this section we prove Theorem 3. We start with the following theorem which will give (1) in Theorem 3 as a corollary.

► **Theorem 4.** *Let  $I \subseteq \{0, 1, \dots, n\}$  be a subset of size  $|I| \leq n/2$ . There does not exist a  $2|I|$ -wise uniform distribution on  $\{0, 1\}^n$  that is supported on  $S := \{x \in \{0, 1\}^n : \sum_i x_i \in I\}$ .*

**Proof.** Suppose there exists such a distribution  $D$ . Consider the  $n$ -variate nonzero real polynomial  $p$  defined by

$$p(x) := \prod_{i \in I} \left(-i + \sum_{j=1}^n x_j\right).$$

Note that  $p(x) = 0$  when  $x \in S$ . And so  $\mathbb{E}[p^2(D)] = 0$  in particular. However, since  $p^2$  has degree at most  $2|I|$ , we have  $\mathbb{E}[p^2(D)] = \mathbb{E}[p^2(U)] > 0$ , where  $U$  is the uniform distribution over  $\{0, 1\}^n$ , a contradiction. ◀

**Proof of (1) in Theorem 3.** When  $I$  corresponds to the mod  $m$  test  $S_m$ ,  $|I| \leq n/m + 1$ . ◀

We now move to (2) in Theorem 3. First we prove a lemma that estimates the sum  $\sum_{x \in S_m} (-1)^{\sum_{i=1}^k x_i}$ . Similar bounds have been established elsewhere, cf. e.g. Theorem 2.9 in [15], but we do not know of a reference with an explicit dependence on  $m$ , which will be used in the next section. (2) follows from bounding above the tail of the Fourier coefficients of the indicator function of  $S_m$ .

► **Lemma 5.** *For any  $1 \leq k \leq n - 1$ ,  $|\sum_{x \in S_m} (-1)^{\sum_{i=1}^k x_i}| \leq 2^n \left(\cos \frac{\pi}{2m}\right)^n$ , while for  $k = 0$   $|\sum_{x \in S_m} (-1)^{\sum_{i=1}^k x_i} - 2^n/m| \leq 2^n \left(\cos \frac{\pi}{2m}\right)^n$ . For odd  $m$  the first bound also holds for  $k = n$ .*

**Proof.** Consider an expansion of

$$p(y) = (1 - y)^k (1 + y)^{n-k}$$

into  $2^n$  terms indexed by  $x \in \{0, 1\}^n$  where  $x_i = 0$  indicates that we take the term 1 from the  $i$ 'th factor. It is easy to see that the coefficient of  $y^d$  is  $\sum_{|x|=d} (-1)^{\sum_{i=1}^k x_i}$ . Denote  $\zeta := e^{2\pi i/m}$  as the  $m$ -th root of unity. Recall the identity

$$\frac{1}{m} \sum_{j=0}^{m-1} \zeta^{jd} = \begin{cases} 1 & \text{if } d \equiv 0 \pmod{m} \\ 0 & \text{otherwise.} \end{cases}$$

Thus the sum we want to bound is equal to

$$\frac{1}{m} \sum_{j=0}^{m-1} p(\zeta^j).$$

Note that  $p(\zeta^0) = p(1) = 0$  for  $k \neq 0$  while for  $k = 0$ ,  $p(\zeta^0) = 2^n$ . For the other terms we have the following bound.

► **Claim 6.** *For  $1 \leq j \leq m - 1$ ,  $|p(\zeta^j)| \leq 2^n \left(\cos \frac{\pi}{2m}\right)^k \left(\cos \frac{\pi}{m}\right)^{n-k}$ .*

**Proof.** As  $|1 + e^{i\theta}| = 2|\cos(\theta/2)|$  and  $|1 - e^{i\theta}| = 2|\sin(\theta/2)|$  we have

$$\begin{aligned} |p(\zeta^j)| &= |1 - \zeta^j|^k |1 + \zeta^j|^{n-k} \\ &= 2^n \left(\sin \frac{j\pi}{m}\right)^k \left(\cos \frac{j\pi}{m}\right)^{n-k} \\ &\leq 2^n \left(\cos \frac{\pi}{2m}\right)^k \left(\cos \frac{\pi}{m}\right)^{n-k}, \end{aligned}$$

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where the last inequality holds for odd  $m$  because (1)  $\sin \frac{j\pi}{m}$  is largest when  $j = \frac{m-1}{2}$  or  $j = \frac{m+1}{2}$ , (2)  $\sin(\frac{\pi}{2} - x) = \cos x$ , and (3)  $\cos \frac{j\pi}{m}$  is largest when  $j = 1$  or  $j = m - 1$ . For even  $m$  the term with  $j = m/2$  is 0, as in this case we are assuming that  $k < n$ , and the bounds for odd  $m$  are valid for the other terms. ◀

Therefore, for  $k \neq 0$  we have

$$\left| \sum_{x \in S_m} (-1)^{\sum_{i=1}^k x_i} \right| = \frac{m-1}{m} \cdot 2^n \left( \cos \frac{\pi}{2m} \right)^k \left( \cos \frac{\pi}{m} \right)^{n-k} \leq 2^n \left( \cos \frac{\pi}{2m} \right)^k \left( \cos \frac{\pi}{m} \right)^{n-k},$$

and we complete the proof using the fact that  $\cos(\pi/m) \leq \cos(\pi/2m)$ . For  $k = 0$  we also need to include the term  $p(1) = 2^n$  which divided by  $m$  gives the term  $2^n/m$ . ◀

► **Remark 7.** Clearly the lemma for  $k = 0$  simply is the well known fact that the cardinality of  $S_m$  is very close to  $2^n/m$ . Equivalently, if  $x$  is uniform in  $\{0, 1\}^n$  then the probability that  $\sum_i x_i \in S_m$  is very close to  $1/m$ . The same holds for the probability that  $\sum_i x_i \equiv c \pmod m$  for any fixed  $c$ . This can be seen by using the polynomial  $y^{-c}p(y)$  in the above proof.

**Proof of (2) in Theorem 3.** Let  $f : \{0, 1\}^n \rightarrow \{0, 1\}$  be the characteristic function of  $S_m$ . We first bound above the nonzero Fourier coefficients of  $f$ . Let  $S = S_m$ . By Lemma 5, we have for any  $\beta$  with  $|\beta| = k > 0$ ,

$$|\hat{f}_\beta| = 2^{-n} \sum_{x \in S} (-1)^{\sum_{i=1}^k x_i} \leq \left( \cos \frac{\pi}{2m} \right)^n \leq 2^{-\alpha n},$$

where  $\alpha = -\ln \cos(\pi/2m)$  depends only on  $m$ . Thus, if  $D$  is  $k$ -wise uniform,

$$\begin{aligned} |\mathbb{E}[f(D)] - \mathbb{E}[f(U)]| &\leq \sum_{|\beta| > k} |\hat{f}_\beta| \cdot |\mathbb{E}_{x \sim D}[(-1)^{\sum x_i \beta_i}]| \\ &\leq \sum_{|\beta| > k} |\hat{f}_\beta| \\ &\leq 2^{-\alpha n} \sum_{t=k+1}^n \binom{n}{t} \\ &= 2^{-\alpha n} \sum_{t=0}^{n-k-1} \binom{n}{t}. \end{aligned}$$

For  $k \geq (1 - \delta)n$ , we have an upper bound of  $2^{n(H(\delta) - \alpha)}$ . Pick  $\delta$  small enough so that  $H(\delta) \leq \alpha/2$ . The result follows by setting  $\gamma := \min\{\alpha/2, \delta\}$ . ◀

Note that the above proof fails when  $m$  is even as we cannot handle the term with  $|\beta| = n$ . Finally, we prove (3) in Theorem 3. We use approximation theory.

**Proof of (3) in Theorem 3.** Let  $f : \{0, 1\}^n \rightarrow \{0, 1\}$  be the characteristic function of  $S_m$ . The proof amounts to exhibiting a real polynomial  $p$  in  $n$  variables of degree  $d = c(n/m)(1/\varepsilon)^2$  such that  $f(x) \leq p(x)$  for every  $x \in \{0, 1\}^n$ , and  $\mathbb{E}[p(U)] \leq \varepsilon$  for  $U$  uniform over  $\{0, 1\}^n$ . To see that this suffices, note that  $\mathbb{E}[p(U)] = \mathbb{E}[p(D)]$  for any distribution  $D$  that is  $d$ -wise uniform. Using this and the fact that  $f$  is non-negative, we can write

$$0 \leq \mathbb{E}[f(U)] \leq \mathbb{E}[p(U)] \leq \varepsilon \quad \text{and} \quad 0 \leq \mathbb{E}[f(D)] \leq \mathbb{E}[p(D)] \leq \varepsilon.$$

Hence,  $|\mathbb{E}[f(U)] - \mathbb{E}[f(D)]| \leq \varepsilon$ . This is the method of sandwiching polynomials from [1].

Let us write  $f = g(\sum_i x_i/n)$ , for  $g: \{0, 1/n, \dots, 1\} \rightarrow \{0, 1\}$ . We exhibit a univariate polynomial  $q$  of degree  $d$  such that  $g(x) \leq q(x)$  for every  $x$ , and the expectation of  $q$  under the binomial distribution is at most  $\varepsilon$ . The polynomial  $p$  is then  $q(\sum_i x_i/n)$ .

Consider the continuous, piecewise linear function  $s: [-1, 1] \rightarrow [0, 1]$  defined as follows. The function is always 0, except at intervals of radius  $a/n$  around the inputs  $x$  where  $g$  equals 1, i.e., inputs  $x$  such that  $nx$  is divisible by  $m$ . In those intervals it goes up and down like a ‘ $\Lambda$ ’, reaching the value of 1 at  $x$ . We set  $a = \varepsilon m/10$ .

By Jackson’s theorem, see e.g. [3, Theorem 7.4] or [5], for a degree  $d = O(n\varepsilon^{-1}a^{-1}) = O(n\varepsilon^{-2}m^{-1})$ , there exists a univariate polynomial  $q'$  of degree  $d$  that approximates  $s$  with pointwise error  $\varepsilon/10$ . Our polynomial  $q$  is defined as  $q := q' + \varepsilon/10$ .

It is clear that  $g(x) \leq q(x)$  for every  $x \in \{0, 1/n, \dots, 1\}$ . It remains to estimate  $E[q(U)]$ .

As  $q'$  is a good approximation of  $s$  we have  $E[q(U)] \leq 2\varepsilon/10 + E[s(U)]$ . We noted in Remark 7 that the remainder modulo  $m$  of  $\sum x_i$  is  $\delta$ -close to uniform for  $\delta = \cos(\pi/2m)^n = e^{-O(n/m^2)}$ . Now the function  $s$ , as a function of  $\sum x_i$ , is a periodic function with period  $m$  and if we feed the uniform distribution over  $\{0, 1/n, \dots, m/n\}$  into  $s$  we have  $E[s] \leq \varepsilon/10$ . It follows that if  $n$  is at least a large constant times  $m^2(\log(1/\varepsilon) + \log m)$ , we have  $E[s(U)] \leq 2\varepsilon/10$  and we conclude that  $E[q(U)] \leq 4\varepsilon/10$ .  $\blacktriangleleft$

### 3 Proof of Theorem 2

In this section we prove Theorem 2. Let  $I$  be a subset of  $\{0, 1, \dots, n-1, n\}$  and  $S \subseteq \{0, 1\}^n$  be the subset of strings whose sum  $\sum_i x_i$  belongs to  $I$ . Let  $U_S$  be the uniform distribution over  $S$ . We are going to construct a  $k$ -wise uniform distribution starting from  $U_S$  and changing the weights of  $k+1$  slices of the Hamming cube. In particular, our distribution will be symmetric. We note that since  $S$  is symmetric, if there is a  $k$ -wise uniform distribution supported on it then by a simple symmetrization argument there must also be a symmetric one.

Let  $\varepsilon_t$  be the bias of a parity of size  $t$  under  $U_S$ , i.e.,  $\varepsilon_t := E_{x \in U_S} [(-1)^{\sum_{i=1}^t x_i}]$ . Note that because we are working with symmetric distributions, all parities of the same size have the same bias. Now let  $\varepsilon(t, \ell)$  be the bias of a parity of size  $t$  over the uniform distribution on strings that sum to  $\ell$ . Note that  $\varepsilon(t, \ell)$  is a scaled version of the Kravchuk polynomial of degree  $t$  in the variable  $\ell$ .

We note that  $\varepsilon_t = \sum_{\ell \in I} \Pr_{x \sim U_S} [\sum_j x_j = \ell] \cdot \varepsilon(t, \ell)$ .

Now let  $a_0 < a_1 < \dots < a_k$  be  $k+1$  points in  $I$  that are closest to  $n/2$  and let  $i^*$  be an index that maximizes  $|a_i - \frac{n}{2}|$ . Finally let  $p_i$  be the probability over  $x$  drawn from  $U_S$  that  $x$  sums to  $a_i$ .

We are going to change the  $p_i$  to  $p_i - \Delta_i$  with the goal of making  $\varepsilon_t$  zero for every  $1 \leq t \leq k$ . The effect of the substitution on  $\varepsilon_t$  is to decrease it by  $\sum_{0 \leq i \leq k} \Delta_i \varepsilon(t, a_i)$ .

Thus our goal is to find  $\Delta_i$ ’s so that

$$\begin{aligned} \sum_{i=0}^k \Delta_i \varepsilon(t, a_i) &= \varepsilon_t, \quad \forall t \in \{1, 2, \dots, k\} \\ \sum_{i=0}^k \Delta_i &= 0, \\ 0 \leq p_i - \Delta_i &\leq 1, \quad \forall i \in \{0, \dots, k\}. \end{aligned}$$

Let  $M$  be the  $(k+1) \times (k+1)$  matrix  $M_{t,i} := \varepsilon(t, a_i)$  where  $t, i \in \{0, \dots, k\}$ . Let  $\Delta :=$

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$(\Delta_0, \dots, \Delta_k)^T$  and  $b := (0, \varepsilon_1, \dots, \varepsilon_k)^T$ . Then the first two conditions form the linear system

$$M\Delta = b.$$

We will show that there is a unique solution  $\Delta$  to this system.

To satisfy the third condition, note that  $p_{i^*}$  is the smallest among all the  $p_i$ 's. It will also be the case that  $p_{i^*} \leq 1/2$ . Thus if  $\|\Delta\|_\infty \leq p_{i^*}$  we will also satisfy the third condition and have a  $k$ -wise uniform distribution supported on  $S$ .

Consider the expression  $n^{-t}(\sum_{j=1}^n (-1)^{x_j})^t$ . If we expand this, cancel factors that appear twice, and collect terms, we can rewrite it as

$$n^{-t} \left( \sum_{j=1}^n (-1)^{x_j} \right)^t = \sum_{r=0}^t \gamma_{t,r} \binom{n}{r}^{-1} \sum_{|\beta|=r} (-1)^{\sum x_i \beta_i},$$

for some choice of non-negative values  $\gamma_{t,r}$ , which by plugging in  $x_1 = x_2 = \dots = x_n = 0$  can be seen to satisfy  $\sum_{r=0}^t \gamma_{t,r} = 1$ .

Let  $\alpha_i := (n - 2a_i)/n$ . Taking expectation in the above equation over all the  $x$ 's with sum equal to  $a_i$  we have for every  $i \in \{0, 1, \dots, k\}$ ,

$$\alpha_i^t = ((n - 2a_i)/n)^t = \sum_{r=0}^t \gamma_{t,r} \binom{n}{r}^{-1} \sum_{|\beta|=r} \mathbb{E}[(-1)^{\sum x_i \beta_i}] = \sum_{r=0}^t \gamma_{t,r} \varepsilon(r, a_i). \quad (\text{A})$$

Let  $M_r$  be the  $r$ -th row of  $M$ . We construct a new matrix  $V$  from  $M$  by applying the following row operations  $R$  to  $M$ : For every  $t$ , set  $V_t = \sum_{r=0}^t \gamma_{t,r} M_r$ . It follows from equation (A) that  $V_{t,i} = \alpha_i^t$ , and so  $V = RM$  is a Vandermonde matrix, which is invertible. Hence,

$$\Delta = V^{-1}Rb$$

is a unique solution.

Therefore it suffices to show that  $\|\Delta\|_\infty \leq p_{i^*}$ . Note that  $\|\Delta\|_\infty \leq \|V^{-1}\|_\infty \|Rb\|_\infty$ , where the  $\infty$  norm of a matrix is the maximum sum of the absolute values along any one row.

Moreover, since  $(Rb)_t = \sum_{r=0}^t \gamma_{t,r} b_r$  and  $\sum_{r=0}^t \gamma_{t,r} = 1$ , we have  $\|Rb\|_\infty \leq \|b\|_\infty$ . Hence, it suffices to bound above  $\|V^{-1}\|_\infty$  and  $\|b\|_\infty$ .

### Roadmap for the following claims

To get an idea of the following claims, consider the case  $m = 3$  and  $k = o(n)$ . We first show in Claim 8 that  $\|V^{-1}\|_\infty \leq 2^{o(n)}$ . Then we find it convenient to bound  $\|b\|_\infty$  and  $p_{i^*}$  multiplied by  $|S|$ . We show that  $|S|p_{i^*} \geq 2^{n(1-o(1))}$  in Claim 9. We note that Claims 8, 9 and 10 hold for any symmetric subset  $S$ . Finally, in Claim 11 we use the definition of  $S$  to obtain bounds on  $a_{i^*}$  and  $b$ , and show that  $|S|\|b\|_\infty \leq (2 - \Omega(1))^n$ . Altogether,

$$\|V^{-1}\|_\infty |S| \|b\|_\infty \leq 2^{o(n)} (2 - \Omega(1))^n \leq 2^{n(1-\Omega(1))} \leq |S|p_{i^*},$$

as desired.

► **Claim 8.**  $\|V^{-1}\|_\infty \leq (k+1) \left(\frac{4en}{k}\right)^k$ .

**Proof.** Since  $V$  is a Vandermonde matrix, we can specify the entries of its inverse explicitly. As shown in e.g. [14] we have

$$V_{i,k-j}^{-1} = (-1)^{k-j} \left( \sum_{\substack{|\beta|=j \\ i \notin \beta}} \alpha^\beta \right) \cdot \left( \prod_{s \neq i} (\alpha_s - \alpha_i)^{-1} \right).$$

We now give an upper bound on each of the factors on the R.H.S.

**Bounding**  $\sum_{|\beta|=j, i \notin \beta} \alpha^\beta$

Since  $|\alpha_i| \leq 1$ , this is bounded by the number of terms,  $\binom{k}{j}$ , and hence by  $2^k$ .

**Bounding**  $\prod_{s \neq i} (\alpha_s - \alpha_i)^{-1}$

Since the difference between every pair of distinct  $a_i, a_j$  is at least 1, we have

$$\prod_{s \neq i} (a_s - a_i) \geq (k/2)!^2$$

when  $k$  is even and is at least  $(\frac{k+1}{2})(\frac{k-1}{2})!$  when  $k$  is odd. By a crude form of Stirling's formula,  $n! \geq (n/e)^n$ , and so we get the lower bound  $(k/2e)^k$  in either case. Hence,

$$\prod_{s \neq i} (\alpha_s - \alpha_i)^{-1} \leq n^k \prod_{s \neq i} (a_s - a_i)^{-1} \leq \left(\frac{2en}{k}\right)^k.$$

Putting the bounds together, we have

$$\|V^{-1}\|_\infty \leq (k+1) \max_{i,j} |V_{i,j}^{-1}| \leq (k+1) \left(\frac{4en}{k}\right)^k. \quad \blacktriangleleft$$

Now we give a lower bound on  $p_{i^*}$ .

► **Claim 9.**  $p_{i^*} |S| \geq \frac{2^{n(1-\alpha_{i^*}^2)}}{n+1}$ .

**Proof.** Using the inequalities  $\binom{n}{i} \geq \frac{2^{nH(i/n)}}{n+1}$  and  $H(\frac{1-\varepsilon}{2}) \geq 1 - \varepsilon^2$ , we have

$$p_{i^*} |S| = \binom{n}{a_{i^*}} \geq \frac{2^{nH(\frac{1-\alpha_{i^*}}{2})}}{n+1} \geq \frac{2^{n(1-\alpha_{i^*}^2)}}{n+1}. \quad \blacktriangleleft$$

Therefore,

$$\frac{p_{i^*} |S|}{\|V^{-1}\|_\infty} \geq \frac{2^{n(1-\alpha_{i^*}^2)}}{(n+1)(k+1)\left(\frac{4en}{k}\right)^k} \geq e^{nf(k,n,a_{i^*})},$$

where

$$f(k, n, a_{i^*}) := \ln 2 \cdot (1 - \alpha_{i^*}^2) - \frac{k}{n} \left( \ln \frac{4en}{k} \right) - o(1).$$

We conclude with the following claim.

► **Claim 10.** *If  $e^{nf(k,n,a_{i^*})} \geq \max_{1 \leq t \leq k} \sum_{x \in S} (-1)^{\sum_{i=1}^t x_i}$ , then there exists a  $k$ -wise uniform distribution supported on  $S$ .*

**Proof.** We just showed

$$\frac{p_{i^*} |S|}{\|V^{-1}\|_\infty} \geq e^{nf(k,n,a_{i^*})} \geq \max_{1 \leq t \leq k} \sum_{x \in S} (-1)^{\sum_{i=1}^t x_i} = \|b\|_\infty |S|.$$

Hence,  $\|\Delta\|_\infty \leq \|V^{-1}\|_\infty \|b\|_\infty \leq p_{i^*}$ . ◀

### 3.1 Zero modulo $m$

We have that  $S_m$  consists of all strings with  $\sum x_i \equiv 0 \pmod{m}$ . It follows that  $|\alpha_{i^*}| \leq (k+1)m/2n$ . We now give an upper bound on  $\|b\|_\infty |S|$ .

► **Claim 11.**  $\|b\|_\infty |S| \leq e^{ng(n,m)}$ , where  $g(n,m) := \ln 2 - \frac{1}{2} \left(\frac{\pi}{2m}\right)^2$ .

**Proof.** Note that  $\|b\|_\infty |S| = \sum_{x \in S} (-1)^{\sum_{i=1}^k x_i}$ . By Lemma 5,

$$\sum_{x \in S} (-1)^{\sum_{i=1}^k x_i} \leq 2^n \left(\cos \frac{\pi}{2m}\right)^n \leq e^{ng(n,m)},$$

where in the last two inequalities we used the fact that  $\ln \cos(x) \leq -\frac{x^2}{2}$  for  $x \in [0, \pi/2)$ . ◀

We are now ready to prove Theorem 2.

**Proof of Theorem 2.** Recall that  $|\alpha_{i^*}| \leq (k+1)m/2n$ . By Claim 11 and Claim 10, it suffices to show that  $f(k, n, a_{i^*}) - g(n, m)$  is positive, where recall

$$\begin{aligned} f(k, n, a_{i^*}) &= \ln 2 \cdot (1 - \alpha_{i^*}^2) - \frac{k}{n} \left(\ln \frac{4en}{k}\right) - o(1) \\ &\geq \ln 2 \cdot \left(1 - \left(\frac{(k+1)m}{2n}\right)^2\right) - \frac{k}{n} \left(\ln \frac{4en}{k}\right) - o(1) \end{aligned}$$

and

$$g(n, m) := \ln 2 - \frac{1}{2} \left(\frac{\pi}{2m}\right)^2.$$

Indeed, we have

$$f(k, n, a_{i^*}) - g(n, m) \geq \frac{1}{2} \left(\frac{\pi}{2m}\right)^2 - \frac{k}{n} \left(\ln \frac{4en}{k}\right) - \ln 2 \cdot \left(\frac{(k+1)m}{2n}\right)^2 - o(1),$$

and choosing  $k = \frac{\varepsilon n}{m^2 \ln m}$  for a sufficiently small  $\varepsilon$  makes this quantity positive. ◀

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